THE UNIFORM CONVERGENCE RATE OF KERNEL DENSITY ESTIMATE

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Abstract

In this paper, we study the uniform convergence rate of kernel density estimate f_n and get optimal uniform rate of convergence without the assumption of compact support for kernel function. It is proved that if the density function f satisfies λ -condition and the kernel function K is λ -good (see section 1), then we have

$$\limsup_{n \to \infty} \left(\frac{n}{\log n} \right)^{\lambda/(1+2\lambda)} \sup_{x \in \mathbb{R}^1} |\hat{f}_n(x) - f(x)| \leq \text{const. a.s.}$$

§1. Introduction

Let X_1, \dots, X_n be a random sample drawn from a population with distribution function F and probability density function f (p.d.f.). Denote the empirical distribution function of X_1, X_2, \dots, X_n by F_n . The kernel estimate of f is difined by

$$\hat{f}_n(x) = \frac{1}{nh_n} \sum_{v=1}^n K\left(\frac{x - X_v}{h_n}\right) = \frac{1}{h_n} \int_{-\infty}^{\infty} K\left(\frac{x - y}{h_n}\right) dF_n(y), \qquad (1)$$

where $h_n > 0$ is a constant depending on n and the kernel K is a p. d. f.

This method was proposed by Rosenblatt [1956] and has aroused the interest of many authors. Since then statisticians have been concerned with the problem of uniform convergence of \hat{f}_n to f. The best result was given by Devroye and Wagner [1980]. In the recent years, the problem about uniform strong convergence rate has attracted much attention of statisticians. Under certain conditions imposed on the density f and kernel K, Schuster [1969] proved that

$$\sup_{x\in \mathbb{R}^1} |\hat{f}_n(x) - f(x)| = O(n^{-\frac{1}{4}+\epsilon}) \quad \text{a. s}$$

for sufficiently large *n*. Under the condition that the *r*-th order derivative of f is uniform bounded on R^1 , Singh [1976] proved that, with a suitable choice of K and h_n , we have

 $\sup_{x\in R^1}|\hat{f}_n(x)-f(x)|=O(n^{-r/(2+2r)}\sqrt{\log\log n})a.s.$

for sufficiently large n. For an m-variate p. d. f. f, Susarla [1981] proved that if

Manuscript received July 12, 1983.

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all the first order partial derivatives of f are uniformly bounded, then

$$\sup_{x \in h^m} |\hat{f}_n(x) - f(x)| = O(n^{-1/(2m+2)} \log \log n) \quad \text{a. s.}$$

for sufficiently large n.

If we suppose that f satisfies δ -order Lipschitz condition $0 < \delta \le 1$ (for $\delta = 1$ is the case above), the convergence rate can reach

$$\sup_{x \in \mathbb{R}^m} |\hat{f}_n(x) - f(x)| = O(n^{-\delta/(2m+2\delta)}\sqrt{\log \log n}) \quad \text{a. s.}$$

Lately, Chen Guijing and Zhao Lincheng have studied the same problem. They improved the convergence rate to

$$\limsup_{n\to\infty} \left(\frac{n}{\log n}\right)^{\lambda/(1+2\lambda)} \sup_{x\in \bar{R}} |\hat{f}_n(x) - f(x)| \leq \text{const.} \quad \text{a.s.} \quad (2)$$

for sufficiently large n under the condition that f satisfies λ -condition (refer to the following definition) and kernel has compact support. In this paper, using a new method, we get the same convergence rate as (2) under weaker condition. Precisely, we omit the assumption that K possesses a compact support.

First, we give two definitions.

We say that p. d. f. f(x) satisfies λ -condition if there exists a real number λ , $0 < \lambda \leq 2$ such that

$$\sup_{x} \left| \int_{x-a}^{x+a} f(t) dt - 2af(x) \right| \leq Ra^{\lambda+1}$$

holds for all $\alpha > 0$, where R is a constant.

It is not difficult to verify that if this satisfies δ^{th} Lipschitz condition, then f(x) satisfies the λ -condition with $\lambda = \delta$; if f'(x) satisfies δ^{th} Lipschitz condition, then f(x) satisfies the λ -condition with $\lambda = 1 + \delta$. We say that the kernel function K(x) is λ -good if it satisfies the following conditions:

(i) K(x) is a symmetric p. d. f. with

$$\int_{-\infty}^{\infty} |x|^{\lambda} K(x) dx < +\infty.$$

(ii) K(x) is bounded and strictly decreasing on $[0, \infty)$ or when K(x) possesses compact support S, it is strictly decreasing on $S \cap [0, \infty)$.

We state the theorems whose proofs are given in section 2 and 3, respectively. **Theorem 1.** Suppose that K(x) is λ -good and f(x) satisfies λ -condition and

$$h_n/(n^{-1}\log n)^{\frac{1}{1+2\lambda}}$$

has a finite positive limit as $n \rightarrow \infty$. Then we have

$$\sup_{x \in U^1} |\hat{f}_n(x) - f(x)| \leq A \left(\frac{\log n}{n}\right)^{\frac{n}{1+2\lambda}} \quad a. s.$$

for sufficiently large n, where A is a constant independent of x and n.

Corollary 1. If K(x) is λ -good with $\lambda = \delta$ and f(x) satisfies δ^{th} order Lipschitz condition, then, with a suitable choice of h_n , we have

$$\sup_{x \in R_1} |\hat{f}_n(x) - f(x)| \leq A (n^{-1} \log n)^{\frac{\delta}{1+2\delta}} \quad \text{a. s.}$$
 (4)

for sufficiently large n.

Corollary 2. If K(x) is λ -good with $\lambda = \delta + 1$ and f'(x), the derivative of f(x), satisfies δ^{th} order Lipschitz condition, then, with a suitable choice of h_n , we have

$$\sup_{x \in R_1} |\hat{f}_n(x) - f(x)| \leq A (n^{-1} \log n)^{\frac{1+\delta}{3+2\delta}} \quad \text{a. s.}$$
 (5)

for sufficiently large n.

In order to prove those corollaries, it is enough to verify that f satisfies λ condition with $\lambda = \delta$, $\lambda = \delta + 1$ respectively. Especially, if f(x) possesses bounded
second order derivative, then f'(x) satisfies Lipschitz condition. From Corollary 2,
it follows that

$$\sup_{x \in R_1} |\hat{f}_n(x) - f(x)| \leq A (n^{-1} \log n)^{2/5}$$
 a.s.

for sufficiently large n.

But even if f(x) is k-times (k>2) differentiable, we could not improve the rate of convergence of $\sup_{x \in R_1} |\hat{f}_n(x) - f(x)|$, where $\hat{f}_n(x)$ is kernel estimate with nonnegative kernel. In other words 2/5 is the highest convergence rate for kernel estimate with nonnegative kernel.

Theorem 2. For any fixed sequence $\{h_n\}$ and $\{c_n\}$ such that

$$\lim_{n\to\infty}h_n=0,\ \lim_{n\to8}c_n=\infty$$

and

$$\lim_{n\to\infty}\frac{\log n}{nh_n}=0$$

there exists uniform continuous density function f such that

$$\sup_{x \in R_1} \left| \hat{f}_n(x) - f(x) \right| \ge \frac{1}{c_n} \qquad \text{a. s.}$$
(6)

holds for sufficiently large n, where $\hat{f}_n(x)$ is a kernel estimate with a suitable choice of kernel.

This theorem means that it is impossible to establish any convergence rate of $\sup_{x} |\hat{f}_{n}(x) - f(x)|$ without some further conditions imposed on f besides of being uniformly continuous.

In the definition of kernel estimate, h_n does not depend on random sample. It is natural to use a function $H_n(X_1, \dots, X_n)$ of the sample instead of h_n . Therefore (1) can be replaced by

$${}^{*}_{f_{n}}(x) = \frac{1}{n} \sum_{i=1}^{n} \frac{1}{H_{n}} K\left(\frac{x - X_{i}}{H_{n}}\right).$$
(7)

Form Theorem 1, it is easy to get the following theorem.

Theorem 3. If K(x) is λ -good and p. d. f. f(x) satisfies λ -condition and

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 $|H_n - (n^{-1}\log n)^{\lambda/(1+2\lambda)}| \leq \alpha (n^{-1}\log n)^{\beta\lambda(1+2\lambda)} \qquad \text{a.s.}$

holds for sufficiently large n, then we have, with probability one,

$$f_n(x) - f(x) \mid \leq A(n^{-1}\log n)^{\frac{n}{1+2\lambda}}$$
 for large n,

where β is a positive number, $0 < \beta \leq 1$.

In this paper we shall not discuss the estimator defined by (7) in detail.

§ 2. The Proof of Theorem 1.

In this section, A, A_1 , A_2 \cdots are all constants. We shall use Devroye's probability inequality in proving theorem and state it in a form that is suitable for our use.

Lemma 1. Suppose that X_1, X_2, \dots, X_n is a random sample drawn from a one-dimensional population with probability distribution F. Denote the empirical distribution of X_1, \dots, X_n by F_n . Suppose $T \subset R^1$, $\mathscr{A}_l = \{ [x-l', x+l'] : x \in T, l' \leq l \}$, $\mathscr{A}_{2l} = \{ [x-2l, x+2l] : x \in R^1 \}$, and

$$\sup_{A \in \mathscr{I}_{21}} F(A) \leqslant B \leqslant 1/4.$$

Then for s > 0 and $n \ge \max(1/B, 8B/s^2)$, we have

$$\begin{split} &P\{\sup_{A \in \mathcal{A}_{l}} | F_{n}(A) - F(A) | \geq s \} \\ & \leq 16n^{2} \exp\{-ns^{2}/(64B + 4s)\} + 8n \exp\{-nB/10\}. \end{split}$$

Proof See [2].

Lemma 2. Let $\{h_n\}$ be decreasing sequence with

$$\lim_{n\to\infty}h_n=0$$

and

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$$\lim_{n \to \infty} \frac{\log n}{n h_n} = 0 \tag{8}$$

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$$\sup_{A \in \mathscr{A}_{l_{n}}} F(A) \leq M_0 h_{n_0} \tag{9}$$

then we have

$$\limsup_{n \to \infty} \left(\frac{n}{h_n \log n} \right)^{\frac{1}{2}} \sup_{\mathscr{A}_{l_n}} |F_n(A) - F(A)| \leq C \text{ a. s.}$$
(10)

where M_0 is any fixed constant and C is a constant.

Proof Choose C such that $C^2 > 8M_0$, and take $B = M_0 h_n$ and $s = A \sqrt{h_n \log n/n}$ in Lemma 1. Then for large n,

$$P\left(\left(\frac{n}{h_n \log n}\right)^{1/2} \sup_{\mathcal{A}_{h_n}} |F_n(A) - F(A)| > O\right)$$

$$\leq 16n^2 \exp\{-n(A^2h_n \log n/n)/(M_0h_n + A\sqrt{h_n \log n/n})\}$$

$$+8n \exp\{-nA\sqrt{h_n \log n/n}/10\}$$

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$$\leq 16n^2 \exp\left(-\frac{A^2}{2M_0}\log n\right) + 8n \exp\{-3\log n\} \leq 24n^2.$$

Hence

$$\sum_{n=1}^{\infty} P\left(\left(\frac{n}{h_n \log n}\right)^{1/2} \sup_{\mathcal{A}_{i_n}} |F_n(A) - F(A)| > O\right) < \infty,$$

and (10) follows from the Borel-Cantelli's lemma.

Lemma 3. Let R(x) be a strictly decreasing continuous function defined on $[0, \infty)$ and r(x) is its inverse function, then we have

$$\int_0^\infty R(x)dx = \int_0^{R_0} r(t)dt$$

where $R_0 = R(0)$.

Now we prove Theorem 1.

Let g(x) be the inverse function of K(x) on $[0, \infty)$. It is easy to verify that $g^{1+\lambda}(y)$ is the inverse function of $K(x^{\frac{1}{1+\lambda}})$. Since

$$\int_0^\infty K(x^{\frac{1}{1+\lambda}})dx = \int_0^\infty K(y)y^{\lambda}(1+\lambda)dy \leq \int_{-\infty}^\infty (1+\lambda)|y|^{\lambda}K(y)dy < +\infty,$$

we have

$$\int_{0}^{K_{0}} g^{1+\lambda}(y) dy = (1+\lambda) \int_{0}^{\infty} y^{\lambda} K(y) dy < +\infty, K_{0} = K(0).$$
(11)

In particular, for $\lambda = 0$, we have

$$\int_{0}^{K_{\circ}} g(t)dt = \int_{0}^{\infty} K(y)dy.$$
⁽¹²⁾

From (12), for any n, we can choose an integer N_n such that

$$\left|\sum_{i=1}^{N_n} \frac{2a_{ni}}{N_n} K_0 - 1\right| \leq h_{n,i}^{\lambda}$$
(13)

where

$$a_{ni} = g\left(\frac{iK_0}{N_n}\right), \quad i = 1, 2, \dots N_n.$$
(14)

From (11), we also have

$$\lim_{n\to\infty}\sum_{i=1}^{N_n}\frac{K_0}{N_n}[a_{ni}]^{1+\lambda}=\int_0^\infty [g(x)]^{1+\lambda}dx<+\infty.$$

Hence

$$\sup_{n} \sum_{i=1}^{N_{n}} \frac{K_{0}}{N_{n}} [a_{ni}]^{1+\lambda} < A_{1}.$$
(15)

Let

$$K_{n}(x) = \sum_{i=1}^{N_{n}} \frac{K_{0}}{N_{n}} I_{[-a_{ni},a_{ni}]}(x) = \sum_{i=1}^{N_{n}} \frac{2K_{0}a_{ni}}{N_{n}} K_{ni}(x), \qquad (16)$$

where

$$K_{ni}(x) = \frac{1}{2a_{ni}} I_{[-a_{ni},a_{ni}]}(x)$$

and $I_{[B]}(x)$ denotes the indicator function of B. It is obvious that

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$$|K_n(x) - K(x)| \leq K_0/N_n.$$
⁽¹⁷⁾

 \mathbf{Let}

$$\overset{*}{f}_{n}(x) = \frac{1}{nh_{n}} \sum_{i=1}^{n} K_{n} \left(\frac{(x-x_{i})}{h_{n}} \right).$$
(18)

From (18), (17) and (1), we have

$$|f_n(x) - f_n(x)| \leq K_0 / N_n h_n.$$
 (19)

Certainly, we can choose N_n such that it satisfies the further condition $K_0/N_n h_n \leq h_n^{\lambda}$.

We can write (18) as

$$f_{n}(x) = \frac{1}{nh_{n}} \sum_{i=1}^{n} \sum_{j=1}^{N_{n}} \frac{2K_{0}a_{nj}}{N_{n}} K_{nj}\left(\frac{(x-x_{i})}{h_{n}}\right) = \sum_{j=1}^{N_{n}} \frac{2K_{0}a_{nj}}{N_{n}} f_{nj}(x),$$

where

$$\overset{*}{f}_{nj}(x) = \frac{1}{nh_n} \sum_{i=1}^n K_{nj} \left(\frac{(x-x_i)}{h_n} \right).$$

Hence

$$\begin{aligned} |\mathring{f}_{n}(x) - f(x)| &= \left| \sum_{j=1}^{N_{n}} \frac{2K_{0}a_{nj}}{N_{n}} \mathring{f}_{nj}(x) - f(x) \right| \\ &\leq \sum_{j=1}^{N_{n}} \frac{2K_{0}a_{nj}}{N_{n}} |\mathring{f}_{nj}(x) - f(x)| + \left| \sum_{j=1}^{N_{n}} \frac{2K_{0}a_{nj}}{N_{n}} - 1 \right| f(x). \end{aligned}$$

From (13), the second term on the right side is less than $A_1h_n^{\lambda}$. Now we estimate the first term. For any j, $1 \le j \le N_n$, we have

$$\begin{vmatrix} {}^{*}_{f_{nj}}(x) - f(x) \end{vmatrix} = \left| \frac{N_{n}(x - a_{nj}h_{n}, x + a_{nj}h_{n})}{2na_{nj}h_{n}} - f(x) \right| \\ \leq \frac{1}{2a_{nj}h_{n}} \left| \frac{N_{n}(x - a_{nj}h_{n}, x + a_{nj}h_{n})}{n} - P_{nj}(x) + \frac{1}{2a_{nj}h_{n}} \right| P_{nj}(x) - 2a_{nj}h_{n}f(x) | \triangleq I_{nj}(x) + J_{nj}(x), \qquad (20)$$

where

$$P_{nj}(x) = \int_{x-a_{nj}h_{n}}^{x+a_{nj}h_{n}} f(t)dt = F(x+a_{nj}h_{n}) - F(x-a_{nj}h_{n})$$

and $N_n(a, b)$ denotes the number of X_i 's among X_1, \dots, X_n which lie in the interval [a, b].

It follows from $\lambda\text{-condition}$ that

$$J_{nj}(x) \leqslant Ra_{nj}^{\lambda}h_{n}^{\lambda}, \forall x, j, n.$$
(21)

From Lemma 2, we have

$$\limsup_{n\to\infty} \left(\frac{n\,h_n}{\log n}\right)^{\frac{1}{2}} \sup_{1\leqslant j\leqslant N_n} \sup_{x\in \mathcal{R}} I_{nj}(x) \leqslant 0 \quad \text{a. s.}$$
(22)

From (15) and (19)-(22), we have

$$\begin{split} \limsup_{n \to \infty} \min\left(\sqrt{\frac{nh_n}{\log n}}, \ h_n^{-\lambda}\right) \sup_{x} |\hat{f}_n(x) - f(x)| \\ \leqslant \limsup_{n \to \infty} \frac{K_0}{N_n} \sum_{j=1}^{N_n} \left(\frac{nh_n}{\log n}\right)^{\frac{1}{2}} \sup_{x} I_{nj}(x) \\ + \limsup_{n \to \infty} h_n^{-\lambda} \left[\sum_{j=1}^{N_n} \frac{2a_{nj}K_0}{N_n} R(a_{nj})^{\lambda} h_n^{\lambda} + A_2 h_n^{\lambda}\right] \leqslant A_3 \quad \text{a.s.,} \quad (23) \end{split}$$

where both A_2 and A_3 are constants.

We choose h_n such that

$$\sqrt{(n^{-1}\log n)h_n^{-1}}=h_n^{\lambda},$$

that is $h_n = (n^{-1} \log n)^{\frac{1}{1+2\lambda}}$, then (23) becomes (3). This completes the proof of Theorem 1.

Remarks.

1. In the process of this proof, we can see that Theorem 1 still holds when K(x) is uniform distribution on [-a, a], where a is a positive constant. Therefore, it is enough to suppose that kernel function K is decreasing on $[0, \infty)$ for Theorem 1.

2. About λ -condition and λ -good. We suppose that the kernel function K(x) is symmetric only because it makes the proof simple and the idea clear. In fact, we can change them as follows.

The p. d. f. f(x) is called satisfying λ -condition if

$$\left|\int_{a-a}^{a+b}f(t)dt-(a+b)f(a)\right| \leq R(b+a)^{\lambda+1}$$

holds for all a>0, b>0 and $x \in R_1$.

The kernel function K(x) is called λ -good if

$$\int_{-\infty}^{\infty} |x|^{\lambda} K(x) dx < +\infty$$

and K(x) decreases on $[0, \infty)$ and increases on $(-\infty, 0]$ or when K(x) possesses a compact support S, K(x) decreases on $[0, \infty) \cap S$ and increases on $(-\infty, 0] \cap S$.

3. It is not difficult to see that the sequence $\{\hat{f}_n\}$ defined by (1) is asymptotically optimal in C. J. Stone's sense^[8].

4. We can replace λ -condition by

$$\sup_{\boldsymbol{x}\in\bar{R}_1}\left|\frac{1}{2a}\int_{\boldsymbol{x}-a}^{\boldsymbol{x}+a}f(t)dt-f(\boldsymbol{x})\right|=O(a^{\lambda}) \text{ as } a\to 0+.$$
(25)

At the same time, we have to suppose that

$$K(x)x^{1+\lambda} \to 0 \text{ as } x \to \infty$$
(26)

in order to ensure that (21) holds. In fact, we choose N_n so large that (13) and $K_0/N_n h_n \leq h_n^{\Lambda}$ hold, that is

$$N_n \geq \frac{K_0}{h_n^{1+\lambda}}.$$

For any $1 \leq j \leq N_n$, we have

$$a_{nj}h_n \leqslant a_{n1}h_n$$
.

Set

$$M_n = g(h_n^{1+\lambda}).$$

Hence

(27)

$$h_n = [K(M_n)]^{\frac{1}{1+\lambda}},$$

It follows from (27), (28) and (14) that

$$a_{n1}h_n = g\left(\frac{K_0}{N_n}\right)h_n \leqslant g(h_n^{1+\lambda})h_n \leqslant M_n [K(M_n)]^{\frac{1}{1+\lambda}}$$
$$\leqslant [K(M_n)M_n^{1+\lambda}]^{\frac{1}{1+\lambda}}0 \quad \text{a.s. } n \to \infty.$$

§ 3. The Proof of Theorem 2.

Since Theorem 2 can be proved in a way similar to [6], we only give an outline.

For fixed constants a>0, b>0 and d with $0 < d \le 2ab$, we choose a function $g(a_{p}, b, d; x)$ defined on $|x| \le b$ satisfying the following conditions:

- 1) $g(a, b, d; 0) = a, g(a, b, d: \pm b) = 0;$
- 2) $0 \le g(a, b, d; x) \le a$ for $|x| \le b$;
- 3) g(a, b, d; x) is continuous on [-b, b];
- 4) $\int_{-b}^{b} g(a, b, d; x) dx = d.$

It is obvious that there exists such function. Let $\{c_n\}$ is any sequnce such that $c_n \rightarrow \infty$. First, we suppose that $\{C_n\}$ satisfies

$$C_n \leqslant C_n^0 = \sqrt{\frac{nh_n}{8\log n}}.$$
(29)

Let

 $a_n = \frac{3}{O_n}, \ b_n = O_n h_n, \ d_n = \frac{1}{2n}, \ n \ge 1$

and

$$e_1 = C_1 h_1, e_n = 2 \sum_{i=1}^{n-1} C_i h_i + C_n h_n, n \ge 2.$$

Let

$$f(x) = \begin{cases} g(a_n, b_n, d_n; x - e_n) \text{ for } |x - e_n| \leq b_n, n = 1, 2, \cdots \\ 0 \text{ for } x < 0. \end{cases}$$

It is not difficult to see that f is uniformly continuous on R_1 and is a p. d. f.

Let

 $\xi_n = N_n (e_n - b_n, e_n + b_n).$

$$P(|\xi_n/n - d_n| \ge h_n/O_n) \le 2 \exp\{-nh_n^2/O_n^2/(2d_n + h_n/O_n)\}$$

$$\le 2 \exp\left(-\frac{nh_n}{4O_n^2}\right) \le 2n^{-2} \text{ (for sufficiently large } n\text{).}$$
(30)

Let

$$T_n = \{w: w = (X_1 \cdots X_n \cdots) \text{ such that } |\xi_n/n - d_n| \ge h_n/O_n\}.$$

(28)

It follows from (30) and Borel-Cantelli Lemma that

$$P(\limsup_{n\to\infty}T_n)=0.$$

We choose the uniform distribution on [-1, 1] or a symmetric p. d. f. possessing compact support as the kernel function, then for any

$$w = (X_1, X_2 \cdots) \in \limsup T_n$$

we have, for large n,

$$\xi_n \leqslant 2h_n/C_n \text{ and } \hat{f}_n(e_n) \leqslant 2/C_n.$$
(31)

It follows from (31) and the definition of f that

$$|\hat{f}_{n}(e_{n}) - f(e_{n})| \ge \frac{3}{O_{n}} - \frac{2}{O_{n}} - \frac{1}{O_{n}}.$$
 (32)

It is obvious that (32) holds for C_n not satisfying (29). This completes the proof.

References

- [1] Parzen, E., On the estimation of a probability density function and the mode, Ann. Math. Statist., 33 (1962), 1065-1076.
- [2] Devroye, L. P., and Wagner, T. J., The Strong Uniform Consistency of Kernel Density Estimation Multivariate Analysis V., (1980), 55-77.
- [3] Schuster, E. F., Ann. Math. Statist., 40 (1969), 1187.
- [4] Singh, R. S., Ann. Statist., (1976), 431.
- [5] Chen Xiru, Uniform convergence rate of kernel density estimate of a probability density function. Scientia Sinica, to appear.
- [6] Chen Xiru, Convergence rate of the nearest Neibor density estimate, Scientia Sinica, (1981), 12.
- [7] Hoeffbing, W. J., Amer. Statist. Assoc., 58 (1963), 13.
- [8] Stone, C. J., Optimal uniform rate of convergence for nonparametric estimates of a density function or its derivatives, Recet Advances in Statistics: Papers Presented in Honor of Herman Chernoff's Sixtieth Birthday, Rizvi, Bustagi, & Siegmund (eds.), Academic Press, New York.

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