

# UNIFORM STRONG CONVERGENCE RATE OF NEAREST NEIGHBOR DENSITY ESTIMATION

YANG ZENHAI (杨振海)

(Beijing Polytechnic University)

ZHAO LINCHENG (赵林城)

(University of Science and Technology of China)

## Abstract

Based on [3] and [4], the authors study strong convergence rate of the  $k_n$ -NN density estimate  $\hat{f}_n(x)$  of the population density  $f(x)$ , proposed in [1].  $f(x) > 0$  and  $f$  satisfies  $\lambda$ -condition at  $x$  ( $0 < \lambda \leq 2$ ), then for properly chosen  $k_n$

$$\limsup_{n \rightarrow \infty} \left( \frac{n}{\log n} \right)^{\lambda/(1+2\lambda)} |\hat{f}_n(x) - f(x)| \leq C \quad a.s.$$

If  $f$  satisfies  $\lambda$ -condition, then for properly chosen  $k_n$

$$\limsup_{n \rightarrow \infty} \left( \frac{n}{\log n} \right)^{\lambda/(1+2\lambda)} \sup_x |\hat{f}_n(x) - f(x)| \leq C \quad a.s.,$$

where  $C$  is a constant. An order to which the convergence rate of  $|\hat{f}_n(x) - f(x)|$  and  $\sup_x |\hat{f}_n(x) - f(x)|$  cannot reach is also proposed.

## §1. Introduction

Let  $X_1, \dots, X_n$  be i.i.d. samples taking values in  $R$  and having distribution function  $F$  and unknown density function  $f$ . A class of estimators of  $f$  proposed by Loftsgarden and Quesenberry<sup>[1]</sup> has the form

$$\hat{f}_n(x) = k/[2na_n(x)], \quad (1)$$

where  $k = k_n$  is a sequence of positive integer chosen in advance and  $a_n(x)$  equal to the distance from  $x$  to the  $k_n$ th nearest of  $X_1, \dots, X_n$ . Call  $\hat{f}_n(x)$  the nearest neighbor estimator of  $f$ . Since then, this estimate has been widely studied. Devroye and Wagner<sup>[2]</sup> showed the uniform strong consistency of  $f$  when  $f$  is uniformly continuous under the conditions  $k/n \rightarrow 0$  and  $\log n/k \rightarrow 0$ . Concerning the uniform strong consistency and strong convergence rates of  $\hat{f}_n$ , the best result so far as we know was given by Chen ([3] & [4]):

- 1) It is impossible to establish any convergence rate of  $\sup_x |\hat{f}_n(x) - f(x)|$  without

some further conditions imposed on  $f$  besides the uniform continuity.

2) If  $f$  satisfies  $\delta$ -Lipschitz condition,  $0 < \delta \leq 1$ , and we choose

$$k = [n^{2\delta/(1+3\delta)}], \quad (2)$$

then

$$\sup_x |\hat{f}_n(x) - f(x)| = O(n^{-\delta/(1+3\delta)} \sqrt{\log n}) \quad a.s. \quad (3)$$

3) For any  $\delta \in (0, 1]$ , there exists a density function  $f$  satisfying  $\delta$ -Lipschitz condition such that

$$\sup_x |\hat{f}_n(x) - f(x)| = o(n^{-\delta/(1+3\delta)}) \quad a.s. \quad (4)$$

does not hold for any possible choice of  $k$ .

Later, Yang Zhenhai<sup>[5]</sup> proved that if  $\delta=1$ , on choosing  $k$  suitably, the right hand side of (3) can be improved to  $O((\log n/n)^{1/4})$ .

This paper is devoted to further study of this problem.

We call  $f$  satisfies  $\lambda$ -condition at  $x$ ,  $\lambda \in (0, 2]$ , if there exists a  $\lambda \in (0, 1]$  such that

$$|f(x) - f(y)| \leq C|y-x|^\lambda, \quad |y-x| \leq h \quad (5)$$

for  $h$  is small enough, or there exists a  $\lambda \in (1, 2]$  such that

$$|f'(x) - f'(y)| \leq C|y-x|^{\lambda-1}, \quad |y-x| \leq h \quad (6)$$

for  $h$  is sufficiently small, where  $C=C(x)$  is a constant depending on  $x$ .

We say that  $f$  satisfies  $\lambda$ -condition,  $\lambda \in (0, 2]$ , if  $\lambda \in (0, 1]$  and  $f$  satisfies  $\lambda$ -Lipschitz condition, or  $\lambda \in (1, 2]$  and  $f'(x)$  is bounded and satisfies  $(\lambda-1)$ -Lipschitz condition.

We shall prove the following theorems:

**Theorem 1.** Suppose that  $f(x) > 0$  and  $f$  satisfies  $\lambda$ -condition at  $x$  for  $\lambda \in (0, 2]$ .

If we choose

$$k = [n^{2\lambda/(1+2\lambda)} (\log n)^{1/(1+2\lambda)}],$$

then

$$\limsup_{n \rightarrow \infty} \left( \frac{n}{\log n} \right)^{\lambda/(1+2\lambda)} |\hat{f}_n(x) - f(x)| \leq C \quad a.s. \quad (7)$$

where  $C$  is a constant depending on  $x$ . If

$$\lim_{h \rightarrow 0} \frac{1}{h^{1+\lambda}} \int_{x-h}^{x+h} (f(t) - f(x)) dt = e_1 \neq 0, \quad (8)$$

then

$$|\hat{f}_n(x) - f(x)| = o(n^{-\lambda/(1+2\lambda)}) \quad a.s. \quad (9)$$

does not hold for any possible choice of  $k = o(n)$ .

**Theorem 2.** Suppose that  $f$  satisfies  $\lambda$ -condition for  $\lambda \in (0, 2]$ . If we choose

$$k = [n^{2\lambda/(1+3\lambda)} (\log n)^{(1+\lambda)/(1+3\lambda)}], \quad (10)$$

then

$$\limsup_{n \rightarrow \infty} (n/\log n)^{\lambda/(1+3\lambda)} \sup_x |\hat{f}_n(x) - f(x)| \leq C \quad a.s., \quad (11)$$

where  $C$  is a constant not depending on  $n$  and  $x$ .

**Theorem 3.** For any  $\lambda \in (0, 2]$ , there exists a density  $f$  satisfying  $\lambda$ -condition such that

$$\sup_x |\hat{f}_n(x) - f(x)| = o(n^{-\lambda(1+2\lambda)}) \quad a.s. \quad (12)$$

does not hold for any possible choice of  $k$ , and there exists a density function with bounded derivatives of any order such that for any choice of  $k$ , it is impossible that

$$\sup_x |\hat{f}_n(x) - f(x)| = o(n^{-2/7}) \quad a.s. \quad (13)$$

## § 2. Strong Convergence Rate of $f_n(x)$

In this section,  $x$  is fixed and  $C, C_1, C_2, \dots$  are all constants not depending on  $n$  (possibly depending on  $x$ ).

**Lemma 1.** Suppose that the random variable  $Y \sim B(n, p)$ , then for any  $\varepsilon > 0$ , we have

$$P(|Y/n - p| \geq \varepsilon) \leq 2 \exp\{-n\varepsilon^2/(2p + \varepsilon)\}.$$

Refer to [6], Theorem 3.

**Lemma 2.** Suppose  $f(x) > 0$  and  $f$  satisfies  $\lambda$ -condition at  $x, \lambda \in (0, 2]$ , we have

1) if  $k \rightarrow \infty$  and  $k = o(n^{2\lambda/(1+2\lambda)})$  as  $n \rightarrow \infty$ , then

$$\sqrt{k} (\hat{f}_n(x) - f(x)) / f(x) \xrightarrow{L} N(0, 1). \quad (14)$$

2) if (8) holds and  $\lim_{n \rightarrow \infty} C_1 n^{2\lambda/(1+2\lambda)} / k = 1$ , then

$$\sqrt{k} (\hat{f}_n(x) - f(x)) / f(x) - C_1^{1/2} e_1 / (2^{\lambda+1} f^{\lambda+1}(x)) \xrightarrow{L} N(0, 1). \quad (15)$$

3) if the conditions in 2) hold and  $k = o(n)$ , then

$$Z_n \triangleq n^{2\lambda/(1+2\lambda)} |\hat{f}_n(x) - f(x)| \xrightarrow{P} 0$$

is not true.

The argument is similar to that of [3], so we omit the proof.

**Proof of Theorem 1** Choose  $k = [n^{\frac{2\lambda}{1+2\lambda}} (\log n)^{\frac{1}{1+2\lambda}}]$ , without loss of generality, we can suppose

$$\begin{aligned} k &= n^{2\lambda/(1+2\lambda)} (\log n)^{1/(1+2\lambda)}, \\ q_n &= (k/n)^{-\lambda} = (n/\log n)^{\lambda/(1+2\lambda)}. \end{aligned} \quad (16)$$

Hence

$$P\{q_n (\hat{f}_n(x) - f(x)) / f(x) \geq 2C_2\} = P\{a_n(x) \leq d_n\}, \quad (17)$$

where

$$d_n = d_n(x) = k / (2nf(x)(1 + 2C_2 q_n^{-1})) = O(k/n). \quad (18)$$

Since  $f$  satisfies  $\lambda$ -condition at  $x$ , it is not difficult to get

$$p_n \triangleq \int_{x-d_n}^{x+d_n} f(t) dt = 2d_n f(x) + \theta_n C_3 d_n^{1+\lambda}, \quad |\theta_n| \leq 1 \quad (19)$$

for  $n$  large enough.

From (16), we can choose sufficiently large  $C_2$  such that

$$p_n = \frac{k}{n} [1 - (2C_2 q_n^{-1} - C_3 \theta_n(k/n)^\lambda (1 + o(1)))] \leq \frac{k}{n} (1 - C_2 q_n^{-1}) \leq k/n, \quad (20)$$

$$k/n - p_n \geq C_2 q_n^{-1} k/n = C_2 (k/n)^{1+\lambda}. \quad (21)$$

Let  $\mu_n(x, d_n)$  be the empirical measure of  $[x - d_n, x + d_n]$ .

We choose  $C_2^2 \geq 12$  and fix it. From (20), (16) and Lemma 1,

$$\begin{aligned} P\{a_n(x) \leq d_n\} &= P\{\mu_n(x, d_n) - p_n \geq k/n - p_n\} \\ &\leq P\{\mu_n(x, d_n) - p_n \\ &\geq C_2 (k/n)^{1+\lambda}\} \\ &\leq 2 \exp\{-nC_2^2 (k/n)^{2+2\lambda} / [2k/n + C_2 (k/n)^{1+\lambda}]\} \\ &\leq 2 \exp(-C_2^2 \log n / 3) < 2n^{-2}. \end{aligned} \quad (22)$$

From (17) and (22)

$$\sum_n P\{q_n(\hat{f}_n(x) - f(x)) \geq 2C_2 f(x)\} < \infty. \quad (23)$$

Hence, from Borel-Cantelli's lemma,  $\limsup_{n \rightarrow \infty} q_n(\hat{f}_n(x) - f(x))$  is bounded above a. s..

In the same way, we can also prove  $\limsup_{n \rightarrow \infty} q_n(\hat{f}_n(x) - f(x))$  is bounded below a. s..

Hence we have proved the first part of the theorem. The rest follows from 3) of Lemma 2.

### § 3. Strong Convergence Rate of $\sup_x |f_n(x) - f(x)|$

The constants  $C, C_1, C_2, \dots, N, N_1, \dots$  in this section are all independent of  $x$  and  $n$ .

**Lemma 3.** Let  $X_1, \dots, X_n$  be i. i. d. samples taken from a one-dimensional population,  $\mu(A)$  and  $\mu_n(A)$  be its probability distribution and empirical measure respectively. Suppose  $T \subset \mathbb{R}$ ,  $\mathcal{A}_l = \{[x - l', x + l'] : x \in T, l' \leq l\}$  and

$$\sup_{\mathcal{A}_l} \mu(A) \leq b \leq 1/4, \quad (24)$$

then, for  $\varepsilon > 0$  and  $n \geq \max(1/b, 8b/\varepsilon^2)$ , we have

$$P\{\sup_{\mathcal{A}_l} |\mu_n(A) - \mu(A)| \geq \varepsilon\} \leq 16n^2 \exp\{-n\varepsilon^2/(64b + 4\varepsilon)\} + 8n \exp\{-nb/10\}. \quad (25)$$

*Proof* See [2].

**Lemma 4.** Use the notations of Lemma 3. Suppose that  $k = o(n)$  and

$$\sup_{\mathcal{A}_l} \mu(A) \leq 10k/n \triangleq b_n \leq 1/4. \quad (26)$$

Then, for  $r > 0$  and  $0 \leq s \leq 1/2$ , there exists  $C_1 > 0$  such that

$$\limsup_{n \rightarrow \infty} \left\{ \left( \frac{n}{k} \right)^{1+r} (\log n)^s \sup_{\mathcal{A}_l} |\mu_n(A) - \mu(A)| \right\} \leq C_1 \text{ a.s.} \quad (27)$$

whenever

$$k \geq \beta n^{2r/(1+2r)} (\log n)^{(1-2s)/(1+2r)}, \quad (\beta > 0 \text{ is a const.}) \quad (28)$$

$n > b_n^{-1}$  and  $n > 8b_n/\varepsilon_n^2$  hold for large  $n$ , from Lemma 3 we have

$$\begin{aligned} P\{\sup_{A_i} |\mu_n(A) - \mu(A)| \geq \varepsilon_n\} \\ \leq 16n^2 \exp\{-nC_1^2(k/n)^{2+2r}(\log n)^{2s}/[640k/n+4\varepsilon_n]\} + 8ne^{-k} \\ \leq 16n^2 \exp\{-C_1^2\beta^{1+2r} \log n/650\} + 8ne^{-k} \leq 16n^{-2} + 8ne^{-k}. \end{aligned} \quad (29)$$

Hence

$$\sum_n P\{(n/k)^{1+r}(\log n)^{-s} \sup_x |\mu_n(A) - \mu(A)| \geq C_1\} < \infty, \quad (30)$$

and from Borel-Cantelli's lemma, (27) is concluded.

**Lemma 5.** Suppose that the density function  $f$  satisfies  $\lambda$ -Lipschitz condition,  $\lambda \in (0, 1]$ . Then for any  $N_1 > 0$  there exists  $N$  depending only upon  $N_1$ , such that for any  $v_n \geq n^{-N_1}$ , the set  $B_n = \{x: f(x) \geq 2v_n\}$  is compact and there exist  $n^N$  closed intervals  $B_{ni} \subset \{f(x) \geq v_n\}$  such that  $L(B_{ni}) \leq n^{-N_1}$ ,  $\omega_f(B_{ni}) < n^{-N_1}$  and  $\bigcup_{i=1}^{n^N} B_{ni} \supset B_n$ , where  $\omega_f(B) = \sup_{x \in B} f(x) - \inf_{x \in B} f(x)$  and  $L$  the Lebesgue measure.

*Proof* It is obvious that  $B_n$  is compact. From the fact that  $f$  satisfies  $\lambda$ -Lipschitz condition, for any  $x \in B_n$  there exists an open interval  $I_x$  containing  $x$  such that  $L(I_x) = C_2 n^{-N_1/\lambda}$  and  $\omega_f(I_x) \leq v_n$ . By Heine-Borel's theorem, we can choose a finite number of closed  $\bar{I}_x$ 's such that their union  $A \supset B_n$  and  $f(x) \geq v_n$  on  $A$ , where  $\bar{I}_x$  denotes the closure of  $I_x$ .

The set  $A$  is a union of  $l_n$  closed intervals without common points, whose lengths are not less than  $C_2 n^{-N_1/\lambda}$ . Hence

$$C_2 n^{-N_1/\lambda} l_n \leq L(A) \leq L\{f(x) \geq n^{-N_1}\} \leq n^{N_1} \int_{\{f(x) \geq n^{-N_1}\}} f(x) dx \leq n^{N_1}.$$

Therefore  $l_n \leq n^{N_1}$ . Also, the length of every interval not exceed  $n^{N_1}$ . Since  $f$  satisfies  $\lambda$ -Lipschitz condition, it is easy to express those closed intervals by a union of some  $B_{ni}$  satisfying the requirements of Lemma 5.

**Lemma 6.** Let  $X_1, \dots, X_n$  be i.i.d. variables with continuous distribution function  $F$ , and denote their empirical distribution function by  $F_n$ , then

$$\limsup_{n \rightarrow \infty} [\sup_x \sqrt{n} |F_n(x) - F(x)| / \sqrt{2 \log \log n}] = 1, \quad a.s. \quad (31)$$

*Proof* See [7].

*Proof of Theorem 2* We can choose

$$k = n^{2\lambda/(1+3\lambda)} (\log n)^{(1+\lambda)/(1+3\lambda)}, \quad (32)$$

$$\theta v_n = (k/n)^{\lambda/(1+\lambda)}, \quad q_n = \rho v_n^{-1} = \rho \theta (n/k)^{\lambda/(1+\lambda)}, \quad (33)$$

where  $0 < \theta < 1$  and  $\rho > 1$  will be chosen later. We can take  $N_1$  so large that  $v_n \geq n^{-N_1}$ , then  $B_n = \{f(x) \geq 2v_n\}$  is covered by the union of  $n^N$   $B_{ni}$ 's satisfying the conditions of Lemma 5 and  $L(B_{ni}) \leq n^{-N_1}$ ,  $\omega_f(B_{ni}) \leq n^{-N_1}$  for each  $B_{ni}$ . Denote

Let  $\mu(x, d)$  and  $\mu_n(x, d)$  be the probability distribution and empirical measure of  $[x-d, x+d]$  respectively. Taking  $s=0$  and  $r=\lambda/(1+\lambda)$  in Lemma 4, we can

assert with probability one that, for  $n$  large enough, the inequality

$$\mu(x, d) - A_n \leq \mu_n(x, d) \leq \mu(x, d) + A_n \quad (34)$$

holds uniformly for all  $x, d$  satisfying

$$\mu(x, 2d) \leq 10k/n, \quad (35)$$

where

$$A_n = 2C_1(\log n/n)^{(1+2\lambda)/(1+3\lambda)}. \quad (36)$$

Now suppose  $x \in B_n^q = \{f(x) < 2v_n\}$ . Since  $f$  satisfies  $\lambda$ -condition, it is easy to get

$$\mu(x, d) = \int_{x-d}^{x+d} f(t) dt \leq 2df(x) + C_3 d^{1+\lambda} \leq 4d\theta^{-1}(k/n)^{\lambda/(1+\lambda)} + C_3 d^{1+\lambda}.$$

Noticing  $k/n = (\log n/n)^{(1+\lambda)/(1+3\lambda)}$ ,  $k/n \geq 2A_n$  for large  $n$ , we can take  $d_n = C_4(k/n)^{1/(1+3\lambda)}$  with suitably chosen  $C_4$  such that

$$\mu(x, d_n) \leq k/2n \leq k/n - A_n, \quad (37)$$

$$\mu(x, 2d_n) \leq 10k/n. \quad (38)$$

From (34)—(38), we have

$$\mu_n(x, d_n) \leq \mu(x, d_n) + A_n \leq k/n,$$

$$a_n(x) \geq d_n = C_4(k/n)^{1/(1+3\lambda)},$$

$$\hat{f}_n(x) = \frac{k}{2n} C_4^{-1}(k/n)^{-1/(1+3\lambda)} \leq C_5(\log n/n)^{\lambda/(1+3\lambda)}, \quad (39)$$

and it follows that

$$\limsup_{n \rightarrow \infty} \{ (n/\log n)^{\lambda/(1+3\lambda)} \sup_{x \in B_n^q} |\hat{f}_n(x) - f(x)| \} \leq C_6, \quad a. s. \quad (40)$$

On the other hand, for fixed  $C_7 > C$ .

$$P\{\sup_{x \in B_n} |\hat{f}_n(x) - f(x)| \geq C_7 q_n^{-1}\} \leq \sum_{i=1}^{n^N} P\{\sup_{x \in B_{ni}} |\hat{f}_n(x) - f(x)| \geq C_7 q_n^{-1}\}. \quad (41)$$

Denote  $m_i = \min_{B_{ni}} f(x)$  and  $M = \sup_x f(x)$ , we have  $v_n \leq m_i \leq M$ . It is easy to see,  $|a_n(x)$

$-a_n(y)| \leq |x-y|$  for any  $x, y$ . Therefore

$$\inf_{B_{ni}} a_n(x) \geq \sup_{B_{ni}} a_n(x) - n^{-N_1}. \quad (42)$$

Obviously

$$\begin{aligned} P\{\sup_{B_{ni}} |\hat{f}_n(x) - f(x)| \geq C_7 q_n^{-1}\} &\leq P\{\sup_{x \in B_{ni}} \hat{f}_n(x) \geq m_i + C_7 q_n^{-1}\} \\ &+ P\{\inf_{x \in B_{ni}} \hat{f}_n(x) \leq m_i + n^{-N_1} - C_7 q_n^{-1}\} \triangleq I_{ni} + J_{ni}. \end{aligned} \quad (43)$$

Now we estimate  $I_{ni}$ . By (42), we have

$$I_{ni} = P\left\{\inf_{B_{ni}} a_n(x) \leq \frac{k}{2nm_i} / (1 + C_7 q_n^{-1} m_i^{-1})\right\} \leq P\{\sup_{B_{ni}} a_n(x) \leq d_n\}, \quad (44)$$

where

$$d_n = \frac{k}{2nm_i} / (1 + C_7 q_n^{-1} m_i^{-1}) + n^{-N_1}. \quad (45)$$

Take  $\rho$  so large that  $C_7/\rho < 1/8$ , then  $C_7 q_n^{-1} m_i^{-1} = C_7 \rho^{-1} v_n/m_i < 1/8$ . Noticing

$1/(1+x) < 1-7x/8$ , for  $0 \leq x < 1/8$ , we have

$$d_n \leq \frac{k}{2nm_i} \left(1 - \frac{7}{8} C_7 q_n^{-1} m_i^{-1}\right) + n^{-N_1}. \quad (46)$$

Fixed  $C_7$  and  $\rho$ , one can choose  $\theta$  so small that, for  $n$  large enough, the following inequalities hold uniformly for all  $i$  and  $w \in B_n$

$$\begin{aligned} \mu(w, d_n) &= \int_{w-d_n}^{w+d_n} f(t) dt \leq 2d_n f(w) + C_3 d_n^{\lambda+1} \leq 2d_n \left( m_i + n^{-N_1} + \frac{1}{2} C_3 d_n^\lambda \right) \\ &\leq \frac{k}{n} \left( 1 - \frac{3}{4} C_7 q_n^{-1} m_i^{-1} \right) \left( 1 + C_8 \left( \frac{k}{n} \right)^\lambda m_i^{1-\lambda} \right) \\ &\leq \frac{k}{n} \left( 1 - \frac{3}{4} C_7 q_n^{-1} m_i^{-1} + C_8 \rho^{\lambda+1} q_n^{-1} m_i^{-1} \right) \leq \frac{k}{n} \left( 1 - \frac{1}{2} C_7 q_n^{-1} m_i^{-1} \right), \end{aligned} \quad (48)$$

$$\mu(w, 2d_n) \leq 10k/n \leq 1/4, \quad (49)$$

$$k/n - \mu(w, d_n) \geq \frac{k}{2n} C_7 q_n^{-1} M^{-1}, \quad (50)$$

$$\begin{aligned} I_{ni} &\leq P \left\{ \sup_{B_{ni}} (\mu_n(w, d_n) - \mu(w, d_n)) \geq \frac{k}{2n} C_7 q_n^{-1} M^{-1} \right\} \\ &\leq 16n^2 \exp \left\{ -nC_7^2 \left( \frac{k}{2n} \right)^2 q_n^{-2} M^{-2} / \left( 640k/n + \frac{2k}{n} C_7 q_n^{-1} M^{-1} \right) \right\} + 8ne^{-k} \\ &\leq 16n^2 \exp \left\{ -\frac{1}{2600} C_7^2 \rho^{-1} \theta^{-1} M^{-2} \log n \right\} + 8ne^{-k} < \frac{1}{2} n^{-N-2}. \end{aligned} \quad (51)$$

The last one is obtained from Lemma 3. The estimate of  $J_{ni}$  is similar to this. Therefore, on account of (41) and (43)

$$\sum_n P q_n \{ \sup_{w \in B_n} |\hat{f}_n(w) - f(w)| \geq C_7 \} < \infty. \quad (52)$$

By Borel-Cantelli's lemma, we have

$$\limsup_{n \rightarrow \infty} \{ q_n \sup_{w \in B_n} |\hat{f}_n(w) - f(w)| \} \leq C_7 \text{ a.s.} \quad (53)$$

Consequently, (11) follows from (40) and (53).

Finally, we give the proof of Theorem 3. For  $\lambda \in (0, 1]$ , The proof is given in [3]; for  $\lambda \in (1, 2]$ , the proof can be given in a similar way by using Lemmas 2 and 4. We omit it here.

To prove (13), construct a density function

$$f(x) = \begin{cases} C_9 x^2 \exp\{1/(x^2-1)\}, & \text{if } |x| < 1, \\ 0, & \text{if } |x| \geq 1. \end{cases} \quad (54)$$

Obviously, it's derivatives of any order are bounded, and  $f(0) = 0$ ,  $f''(0) = 2C_9/e \triangleq b$ . We can suppose  $R$  satisfies one of the following four assumptions (if necessary, we can choose a suitable subsequence instead of the original one).

A.  $k \geq \alpha n$ ,  $0 < \alpha < 1$  is a constant;

B.  $k = o(n)$  and there is a constant  $\beta > 0$  such that  $k \geq \beta n^{4/7}$ ;

C.  $k \rightarrow \infty$  and  $k = o(n^{4/7})$ ;

D.  $k = k_0$ , a constant.

Case A From (54),  $a_n(0) \leq 1$  and  $\hat{f}_n(0) = k/(2na_n(0)) \geq \alpha/2$ . In this case,

$$|\hat{f}_n(0) - f(0)| \stackrel{\text{a.s.}}{=} O(n^{-2/7})$$

is not true.

Case B Denote  $A_n = 2(\log \log n/n)^{1/2}$ . From B,  $A_n \leq k/n$  for large  $n$ . We have with probability one that

$$\mu(d_n) \triangleq \mu[-d_n, d_n] \leq \mu_n[-d_n, d_n] + A_n \triangleq \mu_n(d_n) + A_n \quad (55)$$

for  $n$  large enough and for all  $d_n$ . Taking

$$d_n = \left\{ \frac{4}{b} \left( \frac{k}{n} + A_n \right) \right\}^{1/3},$$

we obtain from  $k = o(n)$  that

$$\mu(d_n) = \int_{-d_n}^{d_n} f(t) dt = \frac{f''(0)}{3} d_n^3 (1 + o(1)) \geq \frac{b}{4} d_n^2 = k/n + A_n,$$

By (55), we have  $\mu(d_n) \geq k/n$ . From the definition of  $a_n(0)$ ,  $a_n(0) \leq d_n \leq C_{10}(k/n)^{1/3}$ .

Therefore we have with probability one that

$$\hat{f}_n(0) \geq \frac{k}{2n} C_{10}^{-1} (k/n)^{-1/3} = C_{11} (k/n)^{2/3} \geq C_{12} n^{-2/7}, \quad (56)$$

so  $|\hat{f}_n(0) - f(0)| \stackrel{a.s.}{=} o(n^{-2/7})$  is not true in case B.

Case C We arbitrarily choose  $x_0 \in (0, 1)$ ,  $f(x_0) \neq 0$ . By 1) of Lemma 2

$$\sqrt{k} (\hat{f}_n(x_0) - f(x_0)) \xrightarrow{L} N(0, 1). \quad (57)$$

From  $k = O(n^{2/7})$ , we see that  $|\hat{f}_n(x_0) - f(x_0)| \stackrel{a.s.}{=} O(n^{-2/7})$  is not true.

Case D For  $x_0$  mentioned in case C

$$P\{n^{2/7}(\hat{f}_n(x_0) - f(x_0)) \leq y\} = P\{a_n(x_0) \geq d_n\}, \quad (58)$$

where

$$d_n = k_0 / \{2nf(x_0)(1 + n^{-2/7}y/f(x_0))\}. \quad (59)$$

Denote  $p_n = \int_{x-d_n}^{x+d_n} f(t) dt$ , we see that  $np_n \rightarrow k_0$  from the continuity of  $f$  at  $x_0$ . Let  $\xi_{n1}, \dots, \xi_{nn}$  be i.i.d. variables with  $P(\xi_{n1} = 1) = p_n$ . According to the Poisson approximation of binomial distribution, we obtain

$$\lim_{n \rightarrow \infty} P\{n^{2/7}(\hat{f}_n(x_0) - f(x_0)) \leq y\} = \lim_{n \rightarrow \infty} P\left\{\sum_{i=1}^n \xi_{ni} \leq k_0\right\} = \sum_{i=0}^{k_0} e^{-k_0} k_0^i / i!. \quad (60)$$

Note that the right hand side of (62) does not depend on  $y$ , so that,

$$|\hat{f}_n(x_0) - f(x_0)| \stackrel{a.s.}{=} O(n^{-2/7})$$

is not true. The proof of Theorem 3 is concluded.

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