WEAK CONVERGENCE OF EMPIRICAL PROCESSES OF SEQUENCES OF STATIONARY RANDOM VARIABLES

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Abstract

In this paper, the author improves Yoshihara's result (J. Multivariate Anal. 8(1978), 584—588) and proves the weak convergence of empirical processes for sequences of ρ -mixing strictly stationary random variable with $\rho(n) = 0(n^{-\frac{1}{2}-\theta})$, $\theta > 0$.

Moreover, the author simplifies the complex proof of weak convergence of empirical processes with random index and gets the corresponding result for α -mixing stationary random variables.

Let $\{\xi_n\}$ be a stationary sequence of random variables defined on a probability space (Ω, \mathcal{F}, P) and let ξ_n have a uniform distribution over [0, 1]. The sequence $\{\xi_n\}$ is said to satisfy the condition of α -mixing, ρ -mixing or φ -mixing, if the following conditions are satisfied respectively:

1.) If for any
$$A \in \mathscr{F}_{-\infty}^0 = \mathscr{F}\{\xi_k: k \leq 0\}, B \in \mathscr{F}_n^\infty = \mathscr{F}\{\xi_k: k \geq n\}, |P(AB) - P(A)P(B)| \leq \alpha(n) \downarrow 0 \quad (n \to \infty);$$

2.) if for any $\mathscr{F}_{-\infty}^0$ -measurable random variable ξ , any \mathscr{F}_n^∞ -measurable random variable η , $E|\xi|^2<\infty$, $E|\eta|^2<\infty$,

$$|E\xi\eta - E\xi E\eta|/\sqrt{\operatorname{Var}\xi\operatorname{Var}\eta} \leqslant \rho(n)\downarrow 0 \quad (n\to\infty);$$

3.) if for any $A \in \mathscr{F}_{-\infty}^0$, $B \in \mathscr{F}_n^\infty$,

$$|P(AB) - P(A)P(B)| \leq \varphi(n)P(A), \varphi(n)\downarrow 0 \quad (n\to\infty).$$

The empirical processes $\{Y_n\}$ are defined as follows:

$$Y_n(t, \omega) = \sqrt{n} (F_n(t, \omega) - t) \quad (0 \le t \le 1), \tag{1}$$

where we denote by $F_n(t, \omega)$ the empirical distribution function of $\xi_1(\omega)$, ..., $\xi_n(\omega)$. Suppose that $Y = \{Y(t, \omega) : 0 \le t \le 1\}$ are real Gaussian processes, with EY(t) = 0 and covariance function

$$EY(s)Y(t) = Eg_s(\xi_0)g_t(\xi_0) + \sum_{k=1}^{\infty} Eg_s(\xi_0)g_t(\xi_k) + \sum_{k=1}^{\infty} Eg_s(\xi_k)g_t(\xi_0), \qquad (2)$$

where $g_t(\alpha) = I_{(0,t)}(\alpha) - t$.

In [1—4], the weak convergence of empirical processes for sequences of stationary random variables has been discussed. Yoshihara proved that Y_n weakly converges to Y in the case of α -mixing and φ -mixing, in [3, 4], when $\alpha(n) = O(n^{-5/2-\theta})$ and $\varphi(n) = O(n^{-1-\theta})$, $\theta > 0$. In this paper we, first of all, weaken the mixing condition and prove the same result in Theorem 1, when $\{\xi_n\}$ is φ -mixing, $\varphi(n) = O(n^{-1/2-\theta})$, $\theta > 0$. Since $\varphi(n) \leq 2\sqrt{\varphi(n)}$, Theorem 1 improves the result in [4].

In [6, 7], the weak convergence of empirical processes for random number of independent random variables has been considered, but the proof is very complex. In Theorem 2, we give a simple proof in the case of dependent variables, that improves the result in [6, 7]

Theorem 1. Let $\{\xi_n\}$ be a stationary ρ -mixing sequence, $\rho(n) = O(n^{-1/2-\theta})$, $\theta > 0$. If the series on the right hand side of (2) are convergent absolutely, then Y_n weakly converges to Y.

Proof It follows from [4, 5] that the finite-dimensional distributions of Y_n converges weakly to those of Y. In order to prove that the $\{Y_n\}$ is tight, from [4], it is sufficient to prove that the following lemma is valid.

Lemma Let
$$\{\xi_n\}$$
 be as in Theorem 1, $\rho(n) = O(n^{-1/2-\theta})$, $\theta > 0$. Put $z_i = I_{[s,t]}(\xi_i) - (t-s)$, $Ez_i = 0$, $Ez_i^2 = \tau > 0$, $S_n = \sum_{k=1}^n z_k$.

Then for any $\varepsilon_0 > 0$, there exist $\theta_1 > 0$, $\theta_2 > 0$ such that

$$P\{|S_n/\sqrt{n}| \geqslant \varepsilon_0\} \leqslant K(n^{-\theta_1}\tau + \tau^{1+\theta_2}), \tag{3}$$

when n is so large and $\tau < \varepsilon_0 / \sqrt{n}$. Here (and below) K is a positive constant (and can assume different values on each of its appearance, even within the same formula).

Proof Let

$$r = [\log_2 n - 1], \ p = 2^{\lfloor (N-1)r/(2N) \rfloor}, \ m = 2^{r - \lfloor r(N-1)/(2N) \rfloor}$$

We may assume that $0 < \theta < 1/2$, and take the nature number $N > 2/(5\theta)$, where $[\alpha]$ denotes the integral part of real number α . Put

$$\begin{split} & \eta_{j} = \sum_{i=1}^{p} z_{jp+i}, \qquad T_{k} = \sum_{j=1}^{k} \eta_{2j-2} \quad (k=1, \ \cdots, \ m), \\ & T'_{m} = \sum_{i=1}^{m} \eta_{2j-1}, \qquad T''_{m} = S_{n} - T_{m} - T'_{m}. \end{split}$$

From the conditions of the lemma, it follows that

$$E\eta_0^2 = E\left(\sum_{i=1}^p z_i\right)^2 \leqslant p\tau + 2\tau((p-1)\rho(1) + \dots + \rho(p-1)) \leqslant Kp^{1+\lambda}\tau,$$

where $\lambda = 1/2 - \theta$. Futhermore

$$\begin{split} E \left| \eta_0 \eta_j \right| \leqslant & E \eta_0^2 \rho(jp) \leqslant K p^{1+\lambda} \tau(jp)^{-(1-\lambda)} \leqslant K p^{2\lambda} j^{-(1-\lambda)} \tau, \\ E T_k^2 \leqslant & K k p^{1+\lambda} \tau + k \sum_{j=1}^{k-1} E \left| \eta_0 \eta_{2j-2} \right| \leqslant & K (k p^{1+\lambda} + k^{1+\lambda} p^{2\lambda}) \tau \leqslant & K k p^{2\lambda} (p^{1-\lambda} \vee k^{\lambda}) \tau, \end{split}$$

where $p^{1-\lambda} \vee k^{\lambda}$ is just the same as max $(\rho^{1-\lambda}, k^{\lambda})$. Since $N > 2/(5\theta) > (1+2\theta)/(5\theta)$, we can take δ such that

We can take
$$\delta$$
 such that $(2(N-1)(1-2\theta)/(N+1-2\theta)) \lor (\{2(N+1)-4(3N-1)\theta\}/\{N-1+2(2N-1)\theta\})$ (4) $<\delta < 2$.

Checking the proof of Lemma 2.1 in [5], we see that Lemma 2.1 is also true for the δ chosen above. It follows from (2.7) and Lemma 2.1 in [5] that we have

$$E|T_{2m}|^{2+\delta} \leq (2+\varepsilon)E|T_m|^{2+\delta} + K(E|T_m|^2)^{1+\delta/2}.$$

There exist positive constants ε and β such that $(2+\varepsilon)$ $2^{-(1+\delta/(2N))} < \beta < 1$. By a discussion as in [4], it follows from Lemma 2.1 of [5] that

$$E|T_{m}|^{2+\delta} \leqslant K(2+\varepsilon)^{r-\left[\frac{N-1}{2N}r\right]-\left[\frac{r\delta}{M}\right]} (E|T(2^{\left[\frac{r\delta}{M}\right]})|^{2})^{1+\delta/2} + K \sum_{i=1}^{r-\left[\frac{N-1}{2N}r\right]-\left[\frac{r\delta}{M}\right]} (2+\varepsilon)^{i-1} (E|T(2^{r-\left[\frac{N-1}{2N}r\right]-i}|^{2})^{1+\delta/2}.$$
(5)

Here M is so large that $(N-1)(1/2+\theta)/(2N) > \delta(1/2-\theta)/M$, which implies $p^{1-\lambda} \ge 2^{\lceil r\delta/M \rceil \lambda}$. Multiplying two hands of the above inequality by $2^{-r(1+\delta/2)}$, we get $2^{-r(1+\delta/2)}E |T_m|^{2+\delta}$

$$\begin{aligned}
& \left\{ K2^{-r\left(1+\frac{\delta}{2}\right)} E\left[T_{m}^{r}\right]^{2+\delta} \\
& \left\{ K2^{-r\left(1+\frac{\delta}{2}\right)+\left(1+-\frac{\delta}{2N}\right)} \left(r-\left[\frac{N-1}{2N}r\right]-\left[\frac{r\delta}{M}\right]\right) + \left[\frac{N-1}{2N}r\right] (1+\lambda)\left(1+\frac{\delta}{2}\right) + \frac{r\delta}{M}\right] (1+\frac{\delta}{2}) \beta^{r-\left[\frac{N-1}{2N}r\right]-\left[\frac{r\delta}{M}\right]} \eta^{1+\frac{\delta}{2}} \\
& + K \sum_{i=1}^{r-\left[-\frac{N-1}{2N}r\right]-\left[\frac{r\delta}{M}\right]} 2^{-r\left(1+\frac{\delta}{2}\right)} (2+s)^{i-1} \eta^{1+\frac{\delta}{2}} \times \\
& \times \left\{ \eta^{1+\lambda} \left(2^{r-\left[-\frac{N-1}{2N}r\right]-i}\right) \vee \eta^{2\lambda} \left(2^{r-\left[-\frac{N-1}{2N}r\right]-i}\right)^{1+\lambda} \right\}^{1+\frac{\delta}{2}}.
\end{aligned} \tag{6}$$

Because of $\tau < \varepsilon_0/\sqrt{n}$, $\tau^{\delta/2} \le K n^{-\delta/4} \le K 2^{-r\delta/4}$. From (4), if M is so large, the powers of 2 in the first term on the right hand side of (6) is

So the first term on the right hand side of (6) does not exceed $Kn^{-\theta_1^{(0)}}\tau$, $\theta_1^{(0)}>0$.

Let us consider the summation of the second term on the right hand side of (6). If $p^{1-\lambda} \ge (2^{r-\lfloor (N-1)r/(2N)\rfloor-i})^{\lambda}$, then the corresponding term in the summation is

$$2^{-r(1+\delta/2)} \{ p^{1+\lambda} 2^{r-\lceil (N-1)r/(2N)\rceil - i} \tau \}^{1+\delta/2} (2+\varepsilon)^{i-1}$$

$$\leq K ((2+\varepsilon))/2^{1+\delta/2})^{i-1} 2^{((N-1)r/(2N))\lambda(1+\delta/2) - r(1-\theta)\delta/4} \tau^{1+\theta\delta/2}.$$
(7)

It follows from (4) that the powers of 2 on the right hand side of (7) is

$$r\left(\frac{N-1}{2N}\left(\frac{1}{2}-\theta\right)\left(1+\frac{\delta}{2}\right)-\frac{1-\theta}{4}\delta\right) = \frac{r}{8N}\left(2(N-1)\left(1-2\theta\right)-(N+1-2\theta)\delta\right)<0.$$

If $p^{1-\lambda} < (2^{r-l(N-1)r/(2N)l-i})$, then the corresponding term in the summation is

$$2^{-r(1+\delta/2)} \{ p^{2\lambda} (2^{r-\lceil (N-1)r/(2N)\rceil-i})^{1+\lambda} \tau \}^{1+\delta/2} (2+s)^{i-1}$$

$$\leq K ((2+s)/2^{1+\alpha})^{i-1} 2^{r(\lambda-(N-1)(1-\lambda)/(2N)-(1-\theta)\delta/(4+2\delta))(1+\delta/2)} \tau^{1+\theta\delta/2},$$
(8)

where $1+\alpha=(1+\lambda)(1+\delta/2)$. It follows from (4) that the powers of 2 on the right hand side of (8) is

$$r(1+\delta/2)(1/2-\theta-(N-1)(1/2+\theta)/(2N)-(1-\theta)\delta/(4+2\delta))$$

$$= (r/8N)\{2(N+1)-4(3N-1)\theta-((N-1)+2(2N-1)\theta)\delta\}<0.$$

Denote $i_0 = \max\{i: p^{1-\lambda} < (2^{r-\lfloor (n-1)r/(2N)\rfloor-i})^{\lambda}\}$. Therefore the second term on the right hand side of (6) does not exceed

$$K \left\{ \sum_{i=1}^{i_0} \left(\frac{2+s}{2^{1+\alpha}} \right)^{i-1} + \sum_{i=i_0}^{r-[(N-1)r/(2N)]-[r\delta/M]} \left(\frac{2+s}{2^{1+\delta/2}} \right)^{i-1} \right\} \tau^{1+\theta\delta/2}$$

$$\leq K \left\{ \sum_{i=1}^{\infty} \left(\left(\frac{2+s}{2^{1+\alpha}} \right)^{i-1} + \left(\frac{2+s}{2^{1+\delta/2}} \right)^{i-1} \right\} \tau^{1+\theta\delta/2} \leqslant K \tau^{1+2\theta_2},$$

where $4\theta_2 = \theta \delta$, the constant K does not depend on r and p. It follows that

$$2^{-r(1+\delta/2)}E|T_m|^{2+\delta} \leq K(n^{-\theta_1^{(0)}}r+r^{1+2\theta_2}).$$

For $T''_m = S_n - T_m - T'_m$, put $0 \le n_1 = n - 2mp < 2^{r+1}$. If $n_1 \le 2p$, since $|z_i| \le 1$, $|T''_m| \le n_1 < 2p$. If $n_1 > 2p$, let us write $r_1 = \lceil \log_2 n_1 - 1 \rceil$, $p_1 = 2^{\lceil (N-1)r_1/(2N) \rceil}$, $m_1 = 2^{r_1}/p_1$, and define $T_k^{(1)}$, $T_{m_1}^{(1)'}$, $T_{m_1}^{(1)'} = T_m'' - T_{m_1}^{(1)} - T_{m_1}^{(1)'}$ as above. By the same argumentation, we have

$$2^{-r_1(1+\delta/2)}E|T_{m_1}^{(1)}|^{2+\delta} \leq K(n^{-\theta_1^{(1)}}\tau + \tau^{1+2\theta_2}).$$

Here the number of terms in $T_{m_1}^{(1)"}$ is $n_2 = n_1 - 2m_1p_1 < 2^{r_1+1} < 2^{r+1}$, therefore $r_1 < r$. For any given n, take this step s times $(s \le r = \lfloor \log_2 n - 1 \rfloor)$ so that $n_i \ge 2p(i=2, \dots, s-1)$, $n_s > 2p$, then $|T_{m_s}^{(s)"}| < 2p < \varepsilon \sqrt{n}/r$ and for $T_{m_s}^{(i)}(i=2, \dots, s-1)$ we have

$$2^{-r_i(1+\delta/2)}E|T_{m_i}^{(i)}|^{2+\delta} \leq K(n^{-\theta_1^{(i)}}\tau + \tau^{1+2\theta_2}).$$

Since $n_1 > n_2 > \cdots > n_{s-1} \ge 2p \ge K n^b(b(>0) \operatorname{doesn't} \operatorname{depend} \operatorname{on} n)$, there exists a $\theta_1(>0)$, which doesn't depend on n, such that $2\theta_1 < \theta_1^{(i)}(i=0, 1, \dots, s-1)$. Put $T_m = T_{m_0}^{(0)}$. It follows that

$$P\{|S_n| \geqslant \varepsilon \sqrt{n}\} \leqslant 2 \sum_{i=0}^{s-1} P\{|T_{m_i}^{(i)}| \geqslant \varepsilon \sqrt{n}/r\} + P\{|T_{m_s}^{(s)''}| \geqslant \varepsilon \sqrt{n}/r\}\}$$
$$\leqslant 2(r/\varepsilon)^{2+\delta} \sum_{i=0}^{s-1} n^{-(1+\delta/2)} E|T_{m_i}^{(i)}|^{2+\delta} \leqslant Kr^{3+\delta} (n^{-2\theta_1}\tau + \tau^{1+2\theta_2}).$$

Note that $r^{3+\delta}/n^{\theta_1}$ and $r^{3+\delta}r^{\theta_2}$ are bounded when n is so large, which implies inequality (3). The proof of the lemma is completed.

Rmark I wish to thank the referee, who has pointed out that the inequality $|T_m''| \leq 2p$ is not always true for any case in [4]. In fact, if $n=2^{r+1}+2^r$, the number of terms in T_m'' is $n_1=n-2mp=2^r$. Then the inequality does not necessarily hold in this case.

Theorem 2. Let $\{\xi_n\}$ be a stationary α -mixing sequence, $\{N_n\}$ be a sequence of random variables with $N_n/n \rightarrow \mu(P_n)$, where μ is a positive random variables. If $Y_n \Rightarrow Y_n$, then $Y_{N_n} \Rightarrow Y_n$.

Proof Write

$$Y_n(t, \omega) = \sqrt{n} (F_n(t, \omega) - t) = n^{-1/2} \sum_{k=1}^n (I(0 \leq \xi_k(\omega) < t) - t).$$

If we can prove that Y_n weakly converges to Y in Renyi-mixing, denoted by $Y_n \Longrightarrow Y$ (R-mixing), then it follows from Theorem 7 in [8], that $Y_{N_n} \Rightarrow Y$.

$$Y'_n(t, \omega) = n^{-1/2} \sum_{k=p_n}^n (I(0 \le \xi_k(\omega) < t) - t),$$

where $p_n \rightarrow \infty$, $p_n / \sqrt{n} \rightarrow 0$. We have

$$\sup_{0 \le t \le 1} |Y_n(t, \omega) - Y'_n(t, \omega)| \le 2p_n / \sqrt{n} \to 0.$$
(9)

Hence $Y'_n \Rightarrow Y$. Let us prove $Y'_n \Rightarrow Y$ (R-mixing), i. e. for any Y-continuity set A and any $E \in \mathscr{F}$,

$$P\{Y_n' \in A, E\} \rightarrow P\{Y \in A\}P\{E\} \quad (n \rightarrow \infty). \tag{10}$$

Let $\mathscr{B}_0 = \bigcup_{n=1}^{\infty} \mathscr{F}(\xi_1, \dots, \xi_n)$ be a class of finite-dimensional sets of the stationary sequence $\{\xi_n\}$. Then for any $E \in \mathcal{B}_0$, there exists a k such that $E \in \mathcal{F}(\xi_1, \dots, \xi_k)$. It follows from α -mixing that we have

$$|P\{Y_n' \in A, E\} - P\{Y_n' \in A\}P\{E\}| \leqslant \alpha(p_n - k) \to 0 \quad (n \to \infty).$$

Because of this, we can prove, as in Theorem 4.5 of [1], that for any F-measurable integrable function g and any Y-continuity set A, we have

$$\int_{(Y_u \in A)} g dP \rightarrow P\{Y \in A\} \int g dP \quad (n \rightarrow \infty).$$

Particularily, if take $g=I_E$, $E\in\mathscr{F}$, then (10) is obtained, i. e. $Y_n'\Rightarrow Y$ (R-mixing). By (9), $\rho(Y'_n, Y_n) \rightarrow 0$. It follows from the lemma of [9] that $Y_n \Rightarrow Y$ (*R*-mixing). The proof is complete.

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